# Statistical Data Analysis 

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## Probability as a Measure of Uncertainty

- Plots and summary statistics are used to learn about the distribution of variables and to investigate their relationships.
- However, we always remain uncertain about the true distributions and relationships in the population since we almost never have access to all of its members.
- Furthermore, our findings based on the observed sample can change if different samples from the population were obtained.
- Therefore, when we generalize our findings from a sample to the whole population, we should explicitly specify the extent of our uncertainty.
- We use probability as a measure of uncertainty.


## Some Commonly Used Genetic Terms

- Genotype

Genetic materials are stored on chromosomes.

- Human somatic cells have two copies of each chromosome
- one inherited from each parent; hence, they are called diploid.

Each pair of similar chromosomes are called homologous chromosomes.
The genotype (i.e., genetic makeup) of an individual for the bi-allelic gene A can take one of the three possible forms:

- Homozygous vs. heterozygous
- The first two genotypes, AA and aa, are called homozygous,
- which means the same version of the allele was inherited from both parents. - That is, both homologous chromosomes have the same allele.
- The last genotype, Aa, is called heterozygous,
- which means different alleles were inherited.


## Some Commonly Used Genetic Terms

- Gene
a segment of double-stranded DNA, which itself is made of a sequence
of four different nucleotides:
- adenine (A), guanine (G), thymine (T), or cytosine (C).
- Single Nucleotide Polymorphisms (SNPs)
- Genetic variation is caused by changes in the DNA sequence of a gene
- SNPs are the most common type of genetic variation.
- SNPs occur when a single nucleotide is replaced by another one. - An example of a SNP would be replacing " G " in the sequence $\{$ TAGCAAT\} by
- Alleles
- alternate forms of a gene
- responsible for variation in phenotypes.
- Phenotypes, in general, are observable traits, such as eye color, disease status,

Phenotypes, in general, are observable traits, such as eye color, disease
and blood pressure, due to genetic factors and/or environmental factors

- In the above example, the alleles could be denoted as T and G.
- We denote the genes with bold face letters (e.g., A) and the two different alleles as capital and small letters (e.g., $A$ and $a$ )


## Some Commonly Used Genetic Terms

- Phenotype
- the set of observable characteristics of an individual resulting from the interaction of its genotype with the environment
- Recessive vs. dominant
- The presence of a specific allele does not always result in its corresponding trait (a characteristic such as eye color).
- Some alleles are recessive,
- producing their trait only when both homologous chromosomes carry that specific variant.
- On the other hand, some alleles are dominant,
- producing their traits when they appear on at least one of the
homologous chromosomes.
- \{For example, suppose that the allele a for gene A is responsible for a specific disease.
Furthermore, assume that a is a recessive allele
- Then, only a person with genotype aa will be affected by the disease.
- Individuals with genotype AA or Aa will not have the disease.\}


## Random phenomena and their sample space

- A phenomenon is called random if its outcome (value) cannot be determined with certainty before it occurs.
- For example, coin tossing and genotypes are random phenomena.
- The collection of all possible outcomes $S$ is called the sample space.

| Coin tossing : | $S=\{H, T\}$, |
| :--- | :--- |
| Die rolling : | $S=\{1,2,3,4,5,6\}$, |
| Bi-allelic gene : | $S=\{A, a\}$, |
| Genotype : | $S=\{A A, A a, a a\}$. |

## Random phenomena and their sample space

- For example, suppose that we suspect that gene $\mathbf{A}$ is related to a specific disease, but genetic variation alone does not determine the disease status.
- Rather, it affects the risk of the disease.
- Further, we suspect that smoking (an environmental factor) is also related to the disease.
- In this case, the random phenomenon we are interested in is the combination of genotype and smoking status
- All possible combinations (i.e., sample space) are identified using the following tree diagram.


## Probability Measure

- To each possible outcome in the sample space, we assign a probability $P$,
- which represents how certain we are about the occurrence
of the corresponding outcome.
- For an outcome $o$, we denote the probability as $P(o)$, - where $0 \leq P(o) \leq 1$.
- The total probability of all outcomes in the sample space is always 1 .
- Coin tossing: $\quad P(H)+P(T)=1$
- Die rolling : $P(1)+P(2)+P(3)+P(4)+P(5)+P(6)=1$
- Therefore, if the outcomes are equally probable,
- the probability of each outcome is $1 / n_{S}$,
- where $n_{S}$ is the number of possible outcomes.


## Random phenomena and their sample space

- The sample space might include an infinite number of possible outcomes.
- For example, the value of blood pressure is random since it cannot be determined with certainty before measuring it.
- The corresponding sample space for blood pressure values is (theoretically) the set of positive real numbers, which is infinite.
- For a complex random phenomenon that is a combination of two or more other random phenomena, it might be easier to view the sample space with tree diagrams.


## Random phenomena and their sample space



## Random events

- An event is a subset of the sample space $S$.
- A possible event for die rolling is
- $E=\{1,3,5\}$.
- This is the event of rolling an odd number.
- For the genotype example,
- $E=\{A A, a a\}$
- This is the event that a person is homozygous.
- An event occurs when any outcome within that event occurs.
- We denote the probability of event $E$ as $P(E)$.
- The probability of an event is the sum of the probabilities for all individual outcomes included in that event.


## Random events - Example 1

- Consider the die rolling example presented in the form of a Venn diagram below.

- All the possible outcomes are contained inside the sample space $S$, which is represented by the rectangle.
- We define two events.
- The event $M$ (shown as a triangle) occurs when the outcome is less than 4.
- The event $N$ (shown as an oval) occurs when the outcome is an odd number.
- In this example, $P(M)=1 / 2$ and $P(N)=1 / 2$


## Random events - Example

- We can define four events as follows:
- The homozygous event : $H M=\{A A, a a\} ;$
- The heterozygous event :
- The no-disease event $H T=\{A a\} ;$
- The disease event $N D=\{A A, A a\} ;$
$D=\{a a\}$
- Assume that the probabilities for different genotypes are - $P(A A)=0.49, P(A a)=0.42$, and $P(a a)=0.09$.
- Then,
$-P(H M)=0.49+0.09=0.58$;
$-P(H T)=0.42$;
$-P(N D)=0.49+0.42=0.91$;
$-P(D)=0.09$.


## Random events - Example 2

- As a running example, we consider a bi-allelic gene $\mathbf{A}$ with two alleles $A$ and $a$.
- We assume that allele $a$ is recessive and causes a specific disease.
- Then only people with the genotype $a a$ have the disease.
- A schematic representation for a bi-allelic gene with a recessive allele $a$ that causes a specific disease.


The shaded area shows the disease event (D)
The unshaded area shows the no-disease even (ND).
The area with shaded border lines shows the homozygous event (HM)
The remaining part of the sample space, which includes the outcome $A a$ only, corresponds to the heterozygous event

## Complement

- For any event $E$, we define its complement, $E^{c}$, as the set of all outcomes that are in the sample space $S$ but not in $E$.
- For the gene-disease example, the complement of the homozygous event $H M=\{A A, a a\}$ is the heterozygous event $\{A a\}$;
- we show this as $H M^{c}=H T$.
- Likewise, the complement of the disease event,
$D=\{a a\}$, is the no-disease event, $N D=\{A A, A a\}$; - we show this as $D^{c}=N D$.
- The probability of the complement event is
- 1 minus the probability of the event:

$$
P\left(E^{c}\right)=1-P(E)
$$

## Complement - example

- For the event that the outcome is an odd number, we have
$-P\left(N^{c}\right)=1-P(N)=1-(1 / 2)=1 / 2$
- equal to the probability that the outcome is an even number.
- In the gene disease example, the probability of the complement of the homozygous event is

$$
-P\left(H M^{c}\right)=1-P(H M)=1-0.58=0.42
$$

- equal to the probability of the heterozygous event $P(H T)=0.42$.
- Likewise, the probability of the complement of the disease event is

$$
-P\left(D^{c}\right)=1-P(D)=1-0.09=0.91
$$

- equal to the probability of the no-disease event, $P(N D)=0.91$


## Complement

- The odds of an event shows how much more certain we are that the event occurs than we are that it does not occur.
- For event $E$, we calculate the odds as follows:
$\frac{P(E)}{P\left(E^{c}\right)}=\frac{P(E)}{1-P(E)}$
- For the gene-disease example, the odds for $N D$ (i.e., not having the disease) are

$$
\frac{P(N D)}{P\left(N D^{c}\right)}=\frac{P(N D)}{1-P(N D)}=\frac{0.91}{1-0.91}=10.11
$$

- Therefore, it is almost 10 times more likely that a person is not affected by the disease than it is for having the disease.
- In this case, we say that the odds for not having the disease are 10 to 1 .


## Union

- For two events $E_{1}$ and $E_{2}$ in a sample space $S$, we define their union $E_{1} \cup E_{2}$ as the set of all outcomes that are at least in one of the events.
- The union $E_{1} \cup E_{2}$ is an event by itself, and it occurs when either $E_{1}$ or $E_{2}$ (or both) occurs.
- For example, the union of the heterozygous event, $H T$, and the disease event, $D$, is - $\{A a\} \cup\{a a\}=\{A a, a a\}$.
- When possible, we can identify the outcomes in the union of the two events and find the probability by adding the probabilities of those outcomes.


## Union

- For the die rolling example (slide 13)

$$
P(M \cup N)=P(\{1,2,3,5\})=\frac{4}{6}=\frac{2}{3}
$$

- Note that in general this is not equal to the sum of the probabilities of the two events:

$$
P(M \cup N) \neq \frac{1}{2}+\frac{1}{2}
$$

- Only under a specific condition, we can write the probability of the union of two events as the sum of their probabilities.
- For the union of the heterozygous event, $H T$, and the disease event, $D$,

$$
P(H T \cup D)=P(\{A a, a a\})=0.42+0.09=0.51
$$

- In this special case, the probability of the union of the two events is equal to the sum of their individual probabilities.


## Intersection

- For two events $E_{1}$ and $E_{2}$ in a sample space $S$, we define their intersection $E_{1} \cap E_{2}$ as the set of outcomes that are in both events.
- The intersection $E_{1} \cap E_{2}$ is an event by itself, and it occurs when both $E_{1}$ and $E_{2}$ occur.
- For example, the intersection of the heterozygous event and the no-disease event is $H M \cap N D=\{A A\}$.
- The intersection of $M$ and $N$ in the dye rolling example (slide 13) is

$$
\begin{aligned}
& M \cap N=\{1,3\} \\
& \text { - In this case, the intersection of the two events includes outcomes } \\
& \text { that are less than } 4 \text { and odd. }
\end{aligned}
$$

- The intersection of the heterozygous event and the nodisease event is $H M \cap N D=\{A A\}$.


## Intersection - Example

- For the die rolling example (slide 13 )

$$
P(M \cap N)=P(\{1,3\})=\frac{2}{6}=\frac{1}{3}
$$

- For the gene-disease example (slide 14)

$$
P(H M \cap N D)=P(A A)=0.49
$$

- Now consider the intersection of the heterozygous event and the disease event.
- There is no common element between $H T$ and $D$.
- Therefore, the intersection is the empty set - $H T \cap D=\{ \}$,
- its probability is
- $P(H T \cap D)=P(\varnothing)=0$.


## Joint vs. marginal probability

- We refer to the probability of the intersection of two events, $P\left(E_{1} \cap E_{2}\right)$, as their joint probability.
- In contrast, we refer to probabilities $P\left(E_{1}\right)$ and $P\left(E_{2}\right)$ as the marginal probabilities of events $E_{1}$ and $E_{2}$.
- For any two events $E_{1}$ and $E_{2}$, we have $-P\left(E_{1} \cup E_{2}\right)=P\left(E_{1}\right)+P\left(E_{2}\right)-P\left(E_{1} \cap E_{2}\right)$.
- That is, the probability of the union $P\left(E_{1} \cap E_{2}\right)$ is the sum of their marginal probabilities minus their joint probability.
- \{The union of the heterozygous and the no-disease events is

$$
\begin{aligned}
-P(H M \cup N D) & =P(H M)+P(N D)-P(H M \cap N D) \\
& =0.58+0.91-0.49=1\}
\end{aligned}
$$

## Disjoint events

- Two events are called disjoint or mutually exclusive if they never occur together:
- if we know that one of them has occurred, we can conclude that the other event has not.
- Disjoint events have no elements (outcomes) in common, and their intersection is the empty set.
- \{For the above example (slide 14), if a person is heterozygous, we know that he does not have the disease

[^0]
## Disjoint events

- For two disjoint events $E_{1}$ and $E_{2}$, the probability of their intersection (i.e., their joint probability) is zero:
$-P\left(E_{1} \cap E_{2}\right)=P(\varphi)=0$
- Therefore, the probability of the union of the two disjoint events is simply the sum of their marginal probabilities:

$$
-P\left(E_{1} \cup E_{2}\right)=P\left(E_{1}\right)+P\left(E_{2}\right)
$$

- In general, if we have multiple disjoint events, $E_{1}$, $E_{2}, \ldots, E_{n}$, then the probability of their union is the sum of the marginal probabilities:

$$
-P\left(E_{1} \cup E_{2} \cup \ldots \cup E_{\mathrm{n}}\right)=P\left(E_{1}\right)+P\left(E_{2}\right)+\ldots+P\left(E_{\mathrm{n}}\right)
$$

## Disjoint events - Example

- The probability of the union of the heterozygous and disease events is
$-P(H T \cup D)=0.42+0.09=0.51$.
- Likewise, when we roll a die, the events $\{1,2\},\{4\}$, and $\{5,6\}$ are disjoint.
- The occurrence of one event prevents the occurrence of the others.
- Therefore, the probability of their union is
$-P(\{1,2\} \cup\{4\} \cup\{5,6\})=1 / 3+1 / 6+1 / 3=5 / 6$
- Now consider the three events $\{1,2,3\},\{4\}$, and $\{5,6\}$.
- These events are disjoint, and their union is the sample space $S$.


## Partition

- When two or more events are disjoint and their union is the sample space $S$,
- we say that the events form a partition of the sample space.
- Two complementary events $E$ and $E^{c}$ always form a partition of the sample space
- since they are disjoint and their union is the sample space.


## Conditional Probability

- Very often, we need to discuss possible changes in the probability of one event based on our knowledge regarding the occurrence of another event.
- The conditional probability, denoted $P\left(E_{1} \mid E_{2}\right)$, is - the probability of event $E_{1}$ given that another event $E_{2}$ has occurred.
- The conditional probability of event $E 1$ given event $E_{2}$ can be calculated as follows: (assuming $P\left(E_{2}\right) \neq 0$ )

$$
P\left(E_{1} \mid E_{2}\right)=\frac{P\left(E_{1} \cap E_{2}\right)}{P\left(E_{2}\right)}
$$

- This is the joint probability of the two events divided by the marginal probability of the event on which we are conditioning .


## Conditional Probability - Example

- Consider the die rolling example (slide 13 ).
- The intersection of the two events is
$-M \cap N=\{1,3\}$
with probability

$$
-P\left(E_{1} \cap E_{2}\right)=2 / 6=1 / 3
$$

- Therefore, the conditional probability of an outcome less than 4 , given that the outcome is an odd number, is

$$
P(M \mid N)=\frac{P(M \cap N)}{P(M)}=\frac{1 / 3}{1 / 2}=\frac{2}{3}
$$

## Conditional Probability - Example

- Consider the gene-disease example (slide 14).
- Suppose we know that a person is homozygous and are interested in the probability that this person has the disease, $P(D \mid H M)$.
- The probability of the intersection of $D$ and $H M$ is $-P(D \cap H M)=P(\{a a\})=0.09$
- Therefore, the conditional probability of having the disease knowing that the genotype is homozygous can be obtained as follows:

$$
P(D \mid H M)=\frac{P(D \cap H M)}{P(H M)}=\frac{0.09}{0.58}=0.16
$$

- In this case, the probability of the disease has increased from $P(D)=0.09$ to $P(D \mid H M)=0.16$.


## Conditional Probability - Example

- Now let us find the conditional probability of not having the disease knowing that the person has a homozygous genotype: $P(N D \mid H M)$.
- The joint probability of $N D$ and $H M$ is $-P(N D \cap H M)=P(\{A A\})=0.49$.
- The conditional probability is therefore

$$
P(N D \mid H M)=\frac{P(N D \cap H M)}{P(H M)}=\frac{0.49}{0.58}=0.84
$$

- The information that the person is homozygous decreases the probability of no disease from its 0.91 to 0.84 .
- Note that the two events $N D$ and $D$ are complementary, and the conditional probability of $N D$ given $H M$ is $-P(N D \mid H M)=1-P(D \mid H M)=1-0.16=0.84$.


## The law of total probability

- By rearranging the equation for conditional probabilities, we have

$$
-P\left(E_{1} \cap E_{2}\right)=P\left(E_{1} \mid E_{2}\right) P\left(E_{2}\right)
$$

- Now suppose that a set of $K$ events $B_{1}, B_{2}, \ldots, B_{\mathrm{K}}$ forms a partition of the sample space.

- Using the above equation, we have $-P(A)=P\left(A \mid B_{1}\right) P\left(B_{1}\right)+\cdots+P\left(A \mid B_{\mathrm{K}}\right) P\left(B_{\mathrm{K}}\right)$
- This is known as the law of total probability


## The law of total probability

- If we know that the event $B_{1}=\{1,2\}$ has occurred, we know for sure that the outcome is less than 4.
- Given $B_{2}=\{3,4\}$, the possible outcomes are now 3 and 4.
- One of two possible outcomes corresponds to the event $M$, that is, the conditional probability of $M$ given $B_{2}$ is $1 / 2$.
- If we know that the event $B_{3}=\{5,6\}$ has occurred, - then the probability that the number is less than 4 is zero:
$P(M \mid B 3)=0$.
- Using the law of total probability, we have

$$
\begin{aligned}
P(M) & =P\left(M \mid B_{1}\right) P\left(B_{1}\right)+P\left(M \mid B_{2}\right) P\left(B_{2}\right)+P\left(M \mid B_{3}\right) P\left(B_{3}\right) \\
& =1 \times \frac{1}{3}+\frac{1}{2} \times \frac{1}{3}+0 \times \frac{1}{3}=\frac{1}{2},
\end{aligned}
$$

which is the same as the probability we found directly based on the outcomes included in $M$.

## Conditional Probability

- In general, all the probability rules we discussed so far apply to conditional probabilities.

- Conditioning on an event only reduces the sample space (e.g., from the large rectangle to the shaded oval in in the figure).
- Within this shrunken sample space, all probability rules are valid.
- For example,

$$
\begin{aligned}
P\left(E_{1}^{c} \mid E_{2}\right) & =1-P\left(E_{1} \mid E_{2}\right) \\
P\left(E_{1} \cup E_{2} \mid E_{3}\right) & =P\left(E_{1} \mid E_{3}\right)+P\left(E_{2} \mid E_{3}\right)-P\left(E_{1} \cap E_{2} \mid E_{3}\right)
\end{aligned}
$$

## The law of total probability

- The law of total probability can be written as

$$
P(A)=\sum_{k=1}^{K} P\left(A \mid B_{k}\right) P\left(B_{k}\right)
$$

where $B_{1}, B_{2}, \ldots, B_{\mathrm{K}}$ form a partition of the sample space, and $A$ is an event in the sample space.

- For die rolling example, consider the three events $-B_{1}=\{1,2\}, B_{2}=\{3,4\}$, and $B_{3}=\{5,6\}$,

$$
\text { - whose probabilities are } P\left(B_{1}\right)=P\left(B_{2}\right)=P\left(B_{3}\right)=1 / 3 \text {. }
$$

- These events form a partition of the sample space.
- The conditional probabilities of $M$ (outcome less than four) given either of these three events are

$$
-P\left(M \mid B_{1}\right)=1, P\left(M \mid B_{2}\right)=1 / 2, P\left(M \mid B_{3}\right)=0
$$

## Independent events

- Two events $E_{1}$ and $E_{2}$ are independent if our knowledge of the occurrence of one event does not change the probability of occurrence of the other event.

$$
\begin{aligned}
& -P\left(E_{1} \mid E_{2}\right)=P\left(E_{1}\right) \\
& -P\left(E_{2} \mid E_{1}\right)=P\left(E_{2}\right)
\end{aligned}
$$

- For example, if a disease is not genetic, knowing a person has a specific genotype (e.g., $A A$ ) does not change the probability of having that disease.


## Independent events

- When two events $E_{1}$ and $E_{2}$ are independent, the probability that $E_{1}$ and $E_{2}$ occur simultaneously, i.e., their joint probability, is the product of their marginal probabilities:
$-P\left(E_{1} \cap E_{2}\right)=P\left(E_{1}\right) \times P\left(E_{2}\right)$
- Therefore, the probability of the union of two independent events is as follows:

$$
-P\left(E_{1} \cup E_{2}\right)=P\left(E_{1}\right)+P\left(E_{2}\right)-P\left(E_{1}\right) \times P\left(E_{2}\right)
$$

- In general, if events $E_{1}, E_{2}, \ldots, E_{\mathrm{n}}$ are independent

$$
-P\left(E_{1} \cap E_{2} \cap \ldots \cap E_{2}\right)=P\left(E_{1}\right) \times P\left(E_{2}\right) \times \ldots \times P\left(E_{\mathrm{n}}\right)
$$

## Disjoint vs Independent events

- Events are disjoined (mutually exclusive) if the occurrence of one event excludes the occurrence of the other(s).
- They cannot happen at the same time.
- For example: when tossing a coin, the result can either be $H$ or $T$ but cannot be both.
- Therefore

$$
-P(H \cap T)=0
$$

$-P(H \cup T)=P(H)+P(T)$
$-P(H \mid T)=0$
$-P\left(H \mid T^{c}\right)=P(H) /\{1-P(T)\}$

## Independent events - Example

- If we toss two fair coins simultaneously, then the probability of observing heads on both coins is - $P\left(H_{1} \cap H_{2}\right)=1 / 2 \times 1 / 2=1 / 4$.
- The probability of the union of two independent events as follows:

$$
-P\left(E_{1} \cup E_{2}\right)=P\left(E_{1}\right)+P\left(E_{2}\right)-P\left(E_{1}\right) \times P\left(E_{2}\right)
$$

- For the above coin tossing example, the probability that at least one of the two coins is heads is

$$
\begin{aligned}
-P\left(H_{1} \cup H_{2}\right) & =1 / 2+1 / 2-1 / 2 \times 1 / 2 \\
& =1-1 / 4=3 / 4=0.75
\end{aligned}
$$

## Disjoint vs Independent events

- Events are independent if the occurrence of one event does not influence (and is not influenced by) the occurrence of the other(s).
- They can happen at the same time.
- For example, when tossing two coins, the result can be $H_{1} H_{2}$, $H_{1} T_{2}, T_{1} H_{2}$, or $T_{1} T_{2}$.
- Considering probability of coming $H_{1} H_{2}$ :
$-P\left(H_{1} \cap H_{2}\right)=P\left(H_{1}\right) P\left(H_{2}\right)$
$-P\left(H_{1} \cup H_{2}\right)=P\left(H_{1}\right)+P\left(H_{2}\right)-P\left(H_{1}\right) P\left(H_{2}\right)$
$-P\left(H_{1} \mid H_{2}\right)=P\left(H_{1}\right)$
$-P\left(H_{1} \mid H_{2}{ }^{c}\right)=P\left(H_{1}\right)$
- This means that disjoint events are not independent, and independent events cannot be disjoint.


## Bayes' theorem

- Sometimes, we know the conditional probability of $E_{1}$ given $E_{2}$, but we are interested in the conditional probability of $E_{2}$ given $E_{1}$.
- For example, suppose that the probability of having lung cancer is $P(C)=0.001$ and that the probability of being a smoker is $P(S M)=0.25$.
- Further, suppose we know that if a person has lung cancer, the probability of being a smoker increases to $P(S M \mid C)=0.40$.
- We are, however, interested in the probability of developing lung cancer if a person is a smoker, $P(C \mid S M)$.


## Bayes' theorem

- In general, for two events $E_{1}$ and $E_{2}$, the following equation shows the relationship between $P\left(E_{2} \mid E_{1}\right)$ and $P\left(E_{1} \mid E_{2}\right)$ :

$$
P\left(E_{2} \mid E_{1}\right)=\frac{P\left(E_{1} \mid E_{2}\right) P\left(E_{2}\right)}{P\left(E_{1}\right)}
$$

- This formula is known as Bayes' theorem or Bayes' rule.
- For the above example,

$$
P(C \mid S M)=\frac{P(S M \mid C) P(C)}{P(S M)}=\frac{0.4 \times 0.001}{0.25}=0.0016
$$

- Therefore, the probability of lung cancer for smokers increases from 0.001 to 0.0016 .


## Bayes' theorem

- Now suppose that a set of K events $B_{1}, B_{2}, \ldots$, $B_{\mathrm{K}}$ forms a partition of the sample space.
- We can write the Bayes' theorem for each of the partitioning events as follows:

$$
P\left(B_{i} \mid A\right)=\frac{P\left(A \mid B_{i}\right) P\left(B_{i}\right)}{P(A)}
$$

- Here, $B_{i}$ is one of the partitioning events, and $A$ is an event in the sample space.


## Application of Bayes' Theorem

- A Venn diagram illustrating a typical medical diagnosis test ("sweat test" to diagnose Cystic Fibrosis)


Here, the following abbreviations are used

- $S$ : sample space,
- H : healthy,
- D : diseased,
- $\mathrm{T}^{-}$: negative test result,
- $\mathrm{T}^{+}$: positive test result.
- The true positive TP : The shaded area to the right of vertical line
- The false positive FP : The shaded area to the left of the vertical line
- The true negative TN : The unshaded area to the left of the vertical
line
- The false negative FN : The unshaded area to the right of the vertical line


## Application of Bayes' Theorem

- Assuming that chromosomes from two parents are transmitted independently, there is the probability $P(D)=0.5 \times 0.5=0.25$ that the child becomes affected (i.e., aa genotype).
- Then, the probability of being healthy is

$$
\text { - } P(H)=1-0.25=0.75
$$

- Assuming that the probability of false positive for the sweat test is $P\left(T^{+} \mid H\right)=0.04$ and the probability of false negative is $P\left(T^{-} \mid D\right)=0.07$
- Because $T^{+}$and $T^{-}$are complementary events, we have

$$
\begin{aligned}
& P\left(T^{-} \mid H\right)=1-P\left(T^{+} \mid H\right)=1-0.04=0.96 \\
& P\left(T^{+} \mid D\right)=1-P\left(T^{-} \mid D\right)=1-0.07=0.93
\end{aligned}
$$

## Bayes' theorem

- Using the law of total probability (slide 34), we have

$$
P(A)=\sum_{k=1}^{K} P\left(A \mid B_{k}\right) P\left(B_{k}\right)
$$

- Therefore, we can write the general form of Bayes' theorem as

$$
P\left(B_{i} \mid A\right)=\frac{P\left(A \mid B_{i}\right) P\left(B_{i}\right)}{\sum_{k=1}^{K} P\left(A \mid B_{k}\right) P\left(B_{k}\right)}
$$

## Application of Bayes' Theorem

- The sweat test is a simple procedure to detect CF by measuring the concentration of salt in a person's sweat. - A high level of salt above a certain cutoff indicates CF.
- The conditional probability of a positive diagnosis for CF patient, $P\left(T^{+} \mid D\right)$, is called the sensitivity of the test.
- The conditional probability of a negative result for a healthy person, $P\left(T^{-} \mid H\right)$, is called the specificity of the test.
- The probability of the CF disease for a child whose parents are both carriers is $P(D)=0.25$.
- Note that the gene causing CF is recessive.
- Therefore, if we denote the allele causing CF as $a$ and the normal allele as $A$, only people with $a a$ genotype have CF.
- People with $A a$ genotype are carriers.
- If both parents are carriers, the chance of transmitting $a$ is 0.5 for each parent


## Application of Bayes' Theorem

- Now we can calculate the updated probability of the disease knowing that the outcome of the test is positive.
- Using the general form of Bayes' theorem, the conditional probability of the disease given a positive test result is

$$
\begin{aligned}
P\left(D \mid T^{+}\right) & =\frac{P\left(T^{+} \mid D\right) P(D)}{P\left(T^{+} \mid D\right) P(D)+P\left(T^{+} \mid H\right) P(H)} \\
& =\frac{0.93 \times 0.25}{0.93 \times 0.25+0.04 \times 0.75}=0.89
\end{aligned}
$$

- Therefore, the positive test result increases the probability of having the disease from $P(D)=0.25$ to $P\left(D \mid T^{+}\right)=0.89$.


## Bayesian Statistics

- In the CF diagnosis example discussed, we assigned the probability of 0.25 to the disease event before seeing any new empirical data.
- This probability is called the prior probability.
- In this case, the prior probability of disease was $P(D)=0.25$.
- After obtaining new evidence, namely positive test results, we updated the probability of the disease from $P(D)$ to $P\left(D \mid T^{+}\right)$.
- We call this updated probability the posterior probability.
- In this case, the posterior probability of the disease was $P\left(D \mid T^{+}\right)=0.89$
- Therefore, based on the test result, we become more certain that the child is affected by the disease.


## Interpretation of Probability as the Relative Frequency

- The random phenomena we have been discussing so far can be observed repeatedly.
- A coin can be tossed or a die can be rolled many times.
- We can observe the genotypes of many people.
- These repeated experiments or observations are called trials.
- For such random phenomena, the probability of an event can be interpreted in terms of the relative frequency.
- The above view of probability is the basis of Frequentist Statistics


## Interpretation of Probability as the Relative Frequency

- As an example, suppose that the probability of genotype $A A$ is $P(A A)=1 / 4$.
- This probability could be interpreted as 1 out of 4 people in the population have genotype $A A$.
- Suppose that we take a simple random sample of size $n$ from the population.
- If the genotype $A A$ is observed $n_{A A}$ times in the sample, the relative frequency of $A A$ in the sample is $n_{A A} / n$.
- If our probability assumption is true (i.e., $P(A A)=1 / 4$ ), this sample relative frequency would be approximately 1/4.
- In this case, as our sample size $n$ increases, the sample relative frequency becomes closer to the probability of $1 / 4$; - that is, it reaches the probability $P(A A)=1 / 4$.


## Interpretation of Probability as the Relative Frequency

- Note that the above interpretation of probability requires two important assumptions.
- We assume that the probability of events does not change from one trial to another.
- For example, the probability of $A A$ must remain $1 / 4$.
- If the population changes as we are sampling people (e.g., genotype $A A$ becomes more prevalent), then the sample relative frequency will not converge to $1 / 4$.
- We also assume that the outcome of one trial does not affect the outcome of another trial.

Interpretation of Probability as the Relative Frequency

- Simulation study of the relative frequency of $A A$ genotype for different sample size values.

- The plot shows how the sample relative frequency of $A A$ genotype approaches the probability $P(A A)$ $=1 / 4$ as the sample size increases.


## Using Tree Diagrams to Obtain Joint Probabilities

- Previously, we used tree diagrams to find the sample space for the combination of two random phenomena.
- Tree diagrams can also be used for calculating their joint probabilities.
- As an example, assume that the alleles on the homologous chromosomes are independent
- i.e., the allele inherited from the mother has no influence on the allele inherited from the father.
- Also assume that for a biallelic gene $\mathbf{A}$, the allele probabilities are $P(A)=0.7$ and $P(a)=0.3$.
- Then to find the genotype probabilities, we can use the tree diagram (shown in next slide).


## Using Tree Diagrams to Obtain Joint Probabilities



- The first set of branches represents possible alleles for one chromosome $\left(C h_{1}\right)$, and the second set represents possible alleles for the other chromosome ( $\mathrm{Ch}_{2}$ ).
- Since these events are independent, knowing the allele on the first chromosome has no influence on the probability of the allele on the second chromosome.


## Using Tree Diagrams to Obtain Joint Probabilities

- The sample space is obtained by following a branch from root to tip:

$$
-S=\left\{A_{1} A_{2}, A_{1} a_{2}, a_{1} A_{2}, a_{1} a_{2}\right\}
$$

- Since these events are independent, their joint probabilities are obtained by multiplying their marginal probabilities:

$$
-P\left(A_{1} A_{2}\right)=0.7 \times 0.7=0.49
$$

- Likewise, the probability of having $a$ on the first chromosome and allele $A$ on the second chromosome is $-P\left(a_{1} A_{2}\right)=0.3 \times 0.7=0.21$
- Following similar approach, we can find the probability of each possible combination of two chromosomes.
- These probabilities are given in the column after the sample space in the figure (previous slide).


## Using Tree Diagrams to Obtain Joint Probabilities

- The labeling of the chromosomes is arbitrary.
- Therefore, we can drop the indices for $A_{1} A_{2}$ and $a_{1} a_{2}$ and write them as genotypes $A A$ and $a a$, respectively.
- The genotype $A a$ can be considered as an event that includes two outcomes,
$-A_{1} a_{2}$ and $a_{1} A_{2}$.
- Therefore, $P(A a)=0.21+0.21=0.42$
- This probability is shown in the last column in the figure (slide 53).


## Using Tree Diagrams to Obtain Joint Probabilities

- Tree diagrams can also be used to find probabilities when the outcomes are not independent.
- Suppose that gene $\mathbf{B}$ in previous example is related to a specific disease, but it is not the only factor to determine the disease status.
- In particular, the probability of having the disease is 0.2 for the $b b$ genotype, whereas this probability is 0.1 for the other two genotypes, $B B$ and $B b$.
- Therefore, the probability of the disease depends on the genotype.


## Using Tree Diagrams to Obtain Joint Probabilities

- The above example can be generalized.
- Assume that the probability of observing the $A$ allele is $P(A)$ $=p$ and the probability of observing the $a$ allele is $P(a)=q$.
- Then the genotype probabilities are
- Homozygous $A A: P\left(A_{1} A_{2}\right)=p \times p=p^{2}$,
- Heterozygous Aa: $P\left(A_{1} a_{2} \cup a_{1} A_{2}\right)=p \times q+q \times p=2 p q$,
- Homozygous aa: $P\left(a_{1} a_{2}\right)=q \times q=q^{2}$.
- Suppose, for example, that the allele probabilities for gene $\mathbf{B}$ are $P(B)=0.8$ and $P(b)=0.2$ and that the alleles on homologous chromosomes are independent (i.e., they are transmitted from parents independently).
- Then the genotype probabilities are
$-P(B B)=0.8=0.64$,
$-P(b b)=0.2=0.04$,
$-P(B b)=2 \times 0.8 \times 0.2=0.32$.

Using Tree Diagrams to Obtain Joint Probabilities


## Using Tree Diagrams to Obtain Joint Probabilities

- Unlike the tree for independent events, the probabilities on the second set of branches depend on the outcomes on the first set of branches.
- As before, we follow the branches from the root to tip and obtain the sample space:
$-S=\{B B-D, B B-H, B b-D, B b-H, b b-D, b b-H\}$.
- To find their probabilities, which are in fact the joint probabilities of genotype and disease status, we multiply the probabilities on the corresponding branches.
- For example, the probability of $B b-D$ is the product of the conditional probability $P(D \mid B b)$ and the marginal probability $P(B b)$ :
$-P(B b-D)=P(B b) P(D \mid B b)=0.32 \times 0.1=0.032$.


[^0]:    - so the two events HT and ND are disjoint. \}

